

# AVERAGED EXTREME REGRESSION QUANTILE

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Consider the linear regression model

$$Y_{ni} = \mathbf{x}_{ni}^\top \boldsymbol{\beta} + e_{ni}, \quad i = 1, \dots, n$$

with observations  $Y_{n1}, \dots, Y_{nn}$ , independent errors  $e_{n1}, \dots, e_{nn}$ , identically distributed according to an unknown distribution function  $F$ ;

$\mathbf{x}_{ni} = (x_{i1}, \dots, x_{ip})^\top$  with  $x_{i1} = 1$ ,  $i = 1, \dots, n$  is the vector of covariates, and  $\boldsymbol{\beta} = (\beta_1, \dots, \beta_p)^\top$  is an unknown parameter. The *maximal regression quantile*  $\hat{\boldsymbol{\beta}}_n(1)$  is a solution of the minimization problem:

$$\min_{\mathbf{b} \in \mathbb{R}^p} \sum_{i=1}^n (Y_i - \mathbf{x}_i^\top \mathbf{b})^+,$$

which can be alternatively described as any solution to the linear program:

$$\min_{\mathbf{b} \in \mathbb{R}^p} \sum_{i=1}^n \mathbf{x}_i^\top \mathbf{b} \quad \text{s.t.} \quad Y_i \leq \mathbf{x}_i^\top \mathbf{b}, \quad i = 1, \dots, n.$$

We are interested in the averaged extreme regression quantile  $\frac{1}{n} \sum_{i=1}^n \mathbf{x}_{ni}^\top \hat{\boldsymbol{\beta}}_n(1)$  and in the relation of  $\frac{1}{n} \sum_{i=1}^n \mathbf{x}_{ni}^\top (\hat{\boldsymbol{\beta}}_n(1) - \boldsymbol{\beta})$  to the extreme quantiles of model errors  $e_{n1}, \dots, e_{nn}$ . Jurečková and Picek (2014) studied the averaged regression  $\alpha$ -quantile  $\hat{\boldsymbol{\beta}}_n(\alpha)$  for  $0 < \alpha < 1$ , and proved that

$$n^{1/2} [\bar{\mathbf{x}}_n^\top (\hat{\boldsymbol{\beta}}_n(\alpha) - \boldsymbol{\beta}) - e_{n:[n\alpha]}] = \mathcal{O}_p(n^{-1/4}) \quad \text{as } n \rightarrow \infty.$$

The situation is different with the averaged extreme regression quantile. Rewrite the regression model in the form  $\mathbf{Y}_n = \mathbf{X}_n \boldsymbol{\beta} + \mathbf{e}_n$ , and consider the following linear programming problem, connected with  $\hat{\boldsymbol{\beta}}_n(1)$ :

$$\begin{aligned} \mathbf{Y}_n^T \mathbf{A} &= \max \\ \mathbf{X}_n^T \mathbf{A} &= \mathbf{X}_n^T \mathbf{1}_n \\ A_i &\geq 0, \quad i = 1, \dots, n \end{aligned}$$

with the optimal base  $\mathbf{X}_1$ , a submatrix of  $\mathbf{X}_n$ , of order  $p \times p$ . Then

$$\bar{\mathbf{x}}_n^\top (\hat{\boldsymbol{\beta}}_n(1) - \boldsymbol{\beta}) = \sum_{i=1}^p \lambda_i e_{n:n-i+1} \quad \text{for any finite } n > p,$$

where  $e_{n:1} \leq \dots \leq e_{n:n}$  are order statistics corresponding to  $e_1, \dots, e_n$  and

$$\boldsymbol{\lambda}^\top = (\lambda_1, \dots, \lambda_p) = \frac{1}{n} \mathbf{1}_n^\top \mathbf{X}_n \mathbf{X}_1^{-1}.$$

The coefficients  $\lambda_i$  are obviously random and satisfy  $\lambda_i > 0$ ,  $i = 1, \dots, p$  and  $\sum_{i=1}^p \lambda_i = 1$ .

**Keywords:**  $\alpha$ -regression quantile, Extreme regression quantile, Averaged extreme regression quantile

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